Are Calculated Betas Good for Anything?

Pablo Fernandez, Professor of Finance IESE Business School, University of Navarra e-mail: <u>fernandezpa@iese.edu</u> and <u>pfernandez@iese.edu</u> Camino del Cerro del Aguila 3. 28023 Madrid, Spain

Previous versions: 2002, 2009, 2013, 2014, 2015. October 18, 2017

We calculate betas of **3,813 companies** using 60 monthly returns each day of December 2001 and January 2002 and show that the median (average) of the maximum beta divided by the minimum beta was 3.07 (15.7). The median of the percentage daily change (in absolute value) of the betas was 20%. Only 2,780 companies (out of 3,813) had all betas positive.

Industry betas are also unstable. On average, the maximum beta of an industry was 2.7 times its minimum beta in December 2001 and January 2002. The median (average) of the percentage daily change (in absolute value) of the industry betas was 7% (16%).

This dispersion of the calculated betas has important implications for the instability of beta-ranked portfolios.

- 1. Historical betas change dramatically from one day to the next
- 2. Implications for making beta-ranked portfolios
- 3. Historical betas depend very much on which index we use to calculate them
- 4. We cannot say that the beta of a company is smaller or bigger than the beta of another
- 5. High-risk companies very often have smaller historical betas than low-risk companies
- 6. Weak correlation between beta and realized return
- 7. About the recommendation of using Industry betas
- 8. Historical betas and the market-to-book ratio.

9. Conclusion

Appendix 1. Short literature review about the CAPM

Appendix 2. Summary statistics of the historical betas of the 30 companies in the Dow Jones Industrial Average Appendix 3. Statistics of the indexes. Historical volatilities, betas with respect to other indexes and correlations

Tables and figures are available in excel format with all calculations in: http://web.iese.edu/PabloFernandez/Book_VaCS/valuation%20CaCS.html

A version in Spanish may be downloaded in: http://ssrn.com/abstract=897700

The beta is one of the most important but elusive parameters in finance. According to the CAPM, it is a measure of the so-called systematic risk. We differentiate the *historical beta* from the *expected beta*, the historical beta being the one we get from the regression of historical data, and the expected beta being the relevant one for estimating the cost of equity (the required return on equity).

Historical betas are used for several purposes:

- To calculate the cost of equity of companies
- To rank assets and portfolios with respect to systematic risk
- To test CAPM and mean-variance efficiency

We argue that historical betas (calculated from historical data) are useless for all three purposes.

The capital asset pricing model (CAPM) defines the required return to equity in the following terms:

$$Ke_i = R_F + E (\beta_i) [E(R_M) - R_F]$$

 R_F = rate of return for risk-free investments (Treasury bonds) E (β_i) = expected equity's beta of company i. E(R_M) = expected market return. [E(R_M) - R_F] = market risk premium

Therefore, given certain values for the equity's beta, the risk-free rate and the market risk premium, it is possible to calculate the required return to equity. The market risk premium is the difference between the *expected* return on the market portfolio and the risk-free rate, which in the context of the CAPM is equal to the incremental return demanded by investors on stocks, above that of risk-free investments.

When estimating betas the standard procedure is to use five years of monthly data and a value-weighted index. This procedure is widely used in academic research and by commercial beta providers such as Merrill Lynch and Ibbotson and Associates. However, different beta sources provide us with different betas, as is shown in **Table 1**. Bruner et al. (1998) also found sizeable differences among beta providers. For their sample the average beta according to Bloomberg was 1.03, whereas according to Value Line it was 1.24.

	AT&T	Boeing	CocaCola	Date
Yahoo	0.61	0.46	0.29	12-febr-03
Multex	0.87	0.66	0.42	12-febr-03
Quicken	1.14	0.66	0.41	12-febr-03
Reuters	0.87	0.68	0.42	12-febr-03
Bloomberg	1.00	1.07	0.64	12-febr-03
Datastream	1.10	1.10	0.37	12-febr-03
Buy&hold	0.84	0.66	0.41	14-febr-03

Table 1. Betas of different companies according to different sources

We show that, in general, it is an enormous error to use the historical beta as a proxy for the expected beta. First, because it is almost impossible to calculate a meaningful beta because historical betas change dramatically from one day to the next; second, because very often we cannot say with a relevant statistical confidence that the beta of one company is smaller or bigger than the beta of another; third, because historical betas do not make much sense in many cases: high-risk companies very often have smaller historical betas than low-risk companies; fourth, because historical betas depend very much on which index we use to calculate them.

Those results are far from being new. For example, Damodaran (2001, page 72) also calculates different betas for Cisco versus the S&P 500:

Beta estimates for Cisco versus the S&P 500. Source: Damodaran (2001, page 72)

Ľ	a commarco for	Cisco versus til	sus the Bar Soo. Source. Damoaaran (2001, p						
		Daily	Weekly	Monthly	Quarterly				
	2 years	1.72	1.74	1.82	2.7				
	5 years	1.63	1.70	1.45	1.78				

Damodaran (1994) also makes this point by calculating the beta of Disney. With daily data, he gets 1.33; 1.38 with weekly data; 1.13 with monthly data; 0.44 with quarterly data; and 0.77 with annual data. With a 3-year period, he gets 1.04; 1.13 with 5 years; and 1.18 with 10 years. Also, the beta depends on the index taken as the benchmark; thus, the beta with respect to the Dow 30 is 0.99; with respect to the S&P 500, it is 1.13, and with respect to the Wilshire 5000, it is 1.05.

We calculate the betas using monthly data every day of the month, not only data of the last day of the month as has usually been done. By doing this, the fact that calculated betas change a lot becomes much clearer. We calculate historical betas for 3,813 companies traded on the New York Stock Exchange (1,462) and the Nasdaq (2,351) **each day** in the 2-month period December 1, 2001 – January 31, 2002 using 5 years of **monthly data¹**. Each day's betas are calculated betas with respect to the S&P 500, using 60 monthly returns. For example, on December 18, 2001, the beta is calculated by running a regression of the 60 monthly returns of the company calculated the 18th of every month, on the 60 monthly returns of the S&P 500 calculated the 18th of every month. We have included only companies that traded in December 1996. Because of this criterion, our sample includes only 450 of the 500 companies that were in the S&P 500 in December 2001.

1. Historical betas change dramatically from one day to the next

The results show that historical betas change dramatically from one day to the next.

Tables 2 and 3 report some statistics about the 62 calculated betas of the 3,813 companies in our sample with respect to the S&P 500 in the two -month period of December 2001 and January 2002. Table 2 shows that only 2,780 companies (73%) had positive betas on the 62 consecutive days. Only 434 companies (11%) had betas bigger than one on the 62 consecutive days. And 2,927 companies (77%) had, in the sample period, a maximum beta more than two times bigger than their minimum beta. Of the 450 companies in the S&P 500, 52% had a maximum beta more than two times bigger than their minimum beta. Of the 30 companies in the DJIA, 40% had a maximum beta more than two times bigger than their minimum beta. Looking at the 101 industry betas, 25% (31%) of the industries had a maximum weighted (unweighted) beta more than two times bigger than their minimum beta.

Table 2. Historical betas of the 3,813 companies in our sample with respect to the S&P 500

Betas are calculated each day in the period 1/12/01-31/1/02 using 5 years of monthly data, i.e. on December 18, 2001, the beta is calculated by running a regression of the 60 monthly returns of the company on the 60 monthly returns of the S&P 500, the returns of each month being calculated on the 18th of each month. The table shows that 2,780 companies (with a combined market capitalization of \$11,956 billion) had positive betas on the 62 days in the period 1/12/01-31/1/02. 434 companies had the 62 betas bigger than 1.0 in the period 1/12/01-31/1/02. For 2,927 companies (77% of the sample), the maximum beta divided by the minimum beta was bigger than 2.

The table also contains the statistics of the 450 companies in our sample that belonged to the S&P 500, and of the 30 companies in the DJIA Index in December 2001.

		Full sample			Ì	S&P 500				ĺ	DJIA			
	Compa	anies	Market	larket Cap.		Companies		Market Cap.			Companies		Marke	t Cap.
	Number	%	\$ bn	%		Number	%	\$ bn	%		Number	%	\$ bn	%
All betas > 0	2,780	73%	11,956	93%		404	90%	8,980	93%		28	93%	3,223	94%
Average beta > 1	1,242	33%	5,758	45%		157	35%	4,273	44%		13	43%	1,839	54%
All betas > 1	434	11%	3,116	24%		71	16%	2,574	27%		7	23%	1,372	40%
Average beta < 0	124	3%	132	1%		10	2%	102	1%		0	0%		0%
All betas < 0	2	0%	97	1%		0	0%		0%		0	0%		0%
Abs(Beta max/beta min) > 2	2,927	77%	6,417	50%		235	52%	4,484	47%		12	40%	1,225	36%
Tota	3,813	100%	12,886	100%		450	100%	9,638	100%		30	100%	3,425	100%

The table contains the same statistics for the betas of 101 industries, both weighted and unweighted.

	Industry weighted betas					Indus	try unv	veighted b	oetas
	Indus	tries	Marke	Market Cap.		Indust	tries	Market	Cap.
	Number	%	\$ bn	%		Number	%	\$ bn	%
All betas > 0	98	97%	12,747	99%		95	94%	12,450	129%
Average beta > 1	20	20%	3,630	28%		26	26%	5,381	56%
All betas > 1	12	12%	2,665	21%		18	18%	3,714	39%
Average beta < 0	1	1%	18	0%		0	0%		0%
All betas < 0	0	0%		0%		0	0%		0%
Abs(Beta max/beta min) > 2	25	25%	1,337	10%		31	31%	3,545	37%
Total	101	100%	12,886	100%		101	100%	9,638	100%

Table 3 shows that the median of the averages of the 62 betas calculated for each company was 0.72 for the 3,813 companies in our full sample, 0.82 for the 450 companies in the S&P 500, and 0.88 for the 30 companies in the DJIA. The median of the difference between the maximum and the minimum of the 62 betas calculated for each company was 0.88 for the 3,813 companies in our full sample, 0.63 for the 450 companies in the S&P 500, and 0.53 for the 30 companies in the DJIA. Note that the difference between the maximum and the minimum is smaller than 4 because we have eliminated 127 companies for which this difference was bigger than 4. The median of the absolute value of the ratio between the maximum and the minimum of the 62 betas calculated for each company was 3.07 for the 3,813 companies in our full sample, 2.11 for the 450 companies in the S&P 500 and 1.77 for the 30 companies in the DJIA. The median of the difference between the maximum and the minimum of the 62 betas calculated for each company, divided by the beta calculated on December 31, 2001, was 1.31 for the 3,813 companies in our full sample, 0.76 for the 450 companies in the S&P 500, and 0.52 for the 30 companies in the DJIA. This statistic was 0.49 for the 101 industry weighted betas, and 0.44 for the 101 industry unweighted betas. From Tables 2 and 3 it is clear that industry betas have less dispersion than company betas. The 30 companies on the DJIA betas have, on average, less dispersion than the 450 companies on the S&P 500, and these have, on average, less dispersion than the 3,813 companies of the full sample. We understand by less dispersion that:

- 1. the median and the average of the difference between the maximum and the minimum of the 62 betas calculated for each company is closer to zero,
- 2. the median and the average of the absolute value of the ratio between the maximum and the minimum of the 62 betas calculated for each company is closer to one, and
- 3. the median and the average of the difference between the maximum and the minimum of the 62 betas calculated for each company, divided by the beta calculated on December 31, 2001, is closer to zero.

Table 3. Summary statistics of the historical betas of the 3,813 companies in our sample with respect to the S&P 500. Betas are calculated each day in the period 1/12/01-31/1/02 using 5 years of monthly data, i.e. on December 18, 2001, the beta is calculated by running a regression of the 60 monthly returns of the company on the 60 monthly returns of the S&P 500. The returns of each month are calculated on the 18th of each month. The table contains the median, the average, the maximum, and the minimum, of:

- Beta average: the average of the 62 betas calculated for each company and industry every day in the period 1/12/01-31/1/02.

- Max - Min: maximum beta minus minimum beta of the 62 betas calculated for each company and industry every day in the period 1/12/01-31/1/02.

- Abs (Max / Min): absolute value of the maximum beta divided by the minimum beta of the 62 betas calculated every day in the period 1/12/01-31/1/02.

- (MAX-Min) / Abs (Beta December 31): maximum beta minus minimum beta of the 62 betas calculated for each company and industry every day in the period 1/12/01-31/1/02 divided by the absolute value of the beta calculated on December 31, 2001.

		Com	pany betas	Indust	ry betas	
		Full sample	S&P 500	Weighted	Unweighted	
	Median	0.72	0.82	0.88	0.84	0.74
Beta average	Average	0.83	0.91	0.94	0.88	0.79
	Maximum	4.46	3.13	1.66	2.78	1.98
	minimum	-1.43	-0.17	0.12	0.06	-0.05
	Maller	0.00	0.72	0.52	0.41	0.25
	Median	0.88	0.63	0.53	0.41	0.35
Max - Min	Average	1.05	0.68	0.53	0.45	0.38
	Maximum	3.99	2.17	0.94	1.02	1.23
	minimum	0.12	0.21	0.27	0.14	0.15
	Median	3.07	2.11	1.77	1.64	1.52
Abs (Max / Min)	Average	15.70	4.76	2.72	2.72	2.57
	Maximum	10116.62	251.72	23.58	35.47	42.03
	minimum	0.10	0.12	1.25	1.12	0.95
	i					
	Median	1.31	0.76	0.52	0.49	0.44
(MAX-Min) /	Average	6.72	2.32	0.76	1.26	0.71
Abs (Beta December 31)	Maximum	2997.32	240.01	3.36	55.62	6.77
	minimum	0.21	0.21	0.24	0.12	0.19

Pablo Fernandez IESE Business School, University of Navarra

Table 4 contains the range of variation of the maximum beta minus the minimum beta of the 62 betas calculated for each company and industry every day in the period 1/12/01-31/1/02. For only seven companies was the difference between the maximum beta and the minimum beta smaller than 0.2. Table 4 also contains the maximum beta minus minimum beta of the 62 betas calculated for each company and industry every day in the period 1/12/01-31/1/02 divided by the absolute value of the beta calculated on December 31, 2001.

Table 4. Historical betas of the 3,813 companies in our sample with respect to the S&P 500.

Betas are calculated each day in the period 1/12/01-31/1/02 using 5 years of monthly data, i.e. on December 18, 2001, the beta is calculated by running a regression of the 60 monthly returns of the company on the 60 monthly returns of the S&P 500. The returns of each month are calculated on the 18th of each month. The table contains the range of variation of:

- the maximum beta minus the minimum beta of the 62 betas calculated for each company and industry every day in the period 1/12/01-31/1/02, and of

- the maximum beta minus the minimum beta of the 62 betas calculated for each company and industry every day in the period 1/12/01-31/1/02, divided by the absolute value of the beta calculated on December 31, 2001.

		Maximum Beta - Minimum Beta									
# companies	5	<u>3 - 3.99</u>	<u>2 - 2.99</u>	<u>1 - 1.99</u>	<u>0.5 - 0.99</u>	<u>0.2 - 0.49</u>	< 0.2	<u>average</u>			
3,813	Full sample	65	268	1,246	1,574	653	7	1.05			
450	S&P 500	0	1	56	250	143	0	0.68			
3,363	Not in the S&P 500	65	267	1,190	1,324	510	7	1.10			
30	DJIA	0	0	6	16	8	0	0.53			
101	Industry weighted	0	0	1	37	59	4	0.45			
101	Industry unweighted	0	0	1	15	77	8	0.38			
		(Ma	ximum Be	ta - Mini	mum Beta)	Abs(Beta D	ecembe	er 31)			
# companies	5	(Ma > 3	ximum Be <u>2 - 2.99</u>	ta - Mini <u>1 - 1.99</u>	mum Beta) 0.5 - 0.99	/Abs(Beta D 0.2 - 0.49	ecembe < 0.2	er 31) <u>average</u>			
# companies	s Full sample	$(Ma) > 3 \\ 800$	ximum Be <u>2 - 2.99</u> 425	ta - Mini <u>1 - 1.99</u> 1,208	mum Beta) 0.5 - 0.99 1,125	/Abs(Beta D 0.2 - 0.49 255	December ≤ 0.2	er 31) <u>average</u> 6.72			
# companies 3,813 450	Full sample S&P 500	(Ma) > 3 800 36	ximum Be <u>2 - 2.99</u> 425 23	ta - Mini <u>1 - 1.99</u> 1,208 95	mum Beta), 0.5 - 0.99 1,125 190	/Abs(Beta D 0.2 - 0.49 255 106	December ≤ 0.2 0	er 31) <u>average</u> 6.72 2.32			
# companies 3,813 450 3,363	Full sample S&P 500 Not in the S&P 500	$(Ma) \ge 3 \\ 800 \\ 36 \\ 764$	ximum Be <u>2 - 2.99</u> <u>425</u> <u>23</u> 402	ta - Mini <u>1 - 1.99</u> 1,208 95 1,113	mum Beta), 0.5 - 0.99 1,125 190 935	/Abs(Beta D 0.2 - 0.49 255 106 149	$\frac{0}{0} = \frac{0}{0}$	er 31) <u>average</u> 6.72 2.32 7.31			
# companies 3,813 450 3,363 30	Full sample S&P 500 Not in the S&P 500 DJIA	$(Ma) \ge 3$ 800 36 764 1	ximum Be <u>2 - 2.99</u> <u>425</u> 23 <u>402</u> 1	ta - Mini <u>1 - 1.99</u> <u>1,208</u> <u>95</u> <u>1,113</u> <u>3</u>	mum Beta), 0.5 - 0.99 1,125 190 935 11	Abs(Beta D 0.2 - 0.49 255 106 149 14	0ecembe <u>< 0.2</u> 0 0 0 0 0 0	er 31) <u>average</u> 6.72 2.32 7.31 0.76			
# companies 3,813 450 3,363 30	Full sample S&P 500 Not in the S&P 500 DJIA Industry weighted	$(Ma) \ge 3 \\ 800 \\ 36 \\ 764 \\ 1 \\ 3$	ximum Be <u>2 - 2.99</u> 425 23 402 1 2	ta - Mini <u>1 - 1.99</u> 1,208 95 1,113 3 12	mum Beta), 0.5 - 0.99 1,125 190 935 11 31	Abs(Beta D 0.2 - 0.49 255 106 149 14 51	$\frac{< 0.2}{0}$	er 31) <u>average</u> 6.72 2.32 7.31 0.76 1.26			

Figure 1 shows the historical betas of AT&T, Boeing and Coca-Cola in the two-month period of December 2001 and January 2002 with respect to the S&P 500

Figure 1. Historical betas of AT&T, Boeing and Coca-Cola. Betas calculated during the two-month period of December 2001 and January 2002 with respect to the S&P 500. Each day, betas are calculated using 5 years of monthly data, i.e. on December 18, 2001, the beta is calculated by running a regression of the 60 monthly returns of the company on the 60 monthly returns of the S&P 500. The returns of each month are calculated on the 18th of the month:



It may be seen that the beta of AT&T varies from 0.32 (January 14, 2002) to 1.02 (December 27, 2001), the beta of Boeing varies from 0.57 (January 30, 2002) to 1.22 (January 20, 2002), and the beta of Coca-Cola varies from 0.55 (December 28, 2001) to 1.11 (January 15, 2002).

A closer look at the data shows that the beta of AT&T is higher than the beta of Boeing 32% of the days, and is higher than the beta of Coca-Cola 50% of the days. The beta of Boeing is higher than the beta of Coca-Cola 76% of the days. AT&T has the maximum beta (of the three companies) 29% of the days and the minimum beta 47% of the days. Boeing has the maximum beta (of the three companies) 58% of the days and the minimum beta 15% of the days. Coca-Cola has the maximum beta (of the three companies) 13% of the days and the minimum beta 38% of the days.

Figure 2 shows the historical betas of Procter & Gamble, Philip Morris and Merck in the two-month period of December 2001 and January 2002 with respect to the S&P 500.

Figure 3 contains the historical betas of AT&T, calculated every day during the period between January 1997 and May 2002. It also contains the historical betas of AT&T, but calculated only the last day of each month.

These three tables and three figures are evidence enough to conclude that calculated betas are very unstable.

Figure 2. Historical betas of Procter and Gamble, Philip Morris and Merck. Betas calculated during the two-month period of December 2001 and January 2002 with respect to the S&P 500.



Figure 3. Historical monthly betas of AT&T. Betas calculated during the 53-month period between January 1997 and May 2002 with respect to the S&P 500. Each day, betas are calculated using 5 years of monthly data, i.e. on December 18, 2001, the beta is calculated by running a regression of the 60 monthly returns of the company on the 60 monthly returns of the S&P 500. The returns of each month are calculated the 18th of the month.



Table 5 contains some statistics of the correlation, and of the company volatility divided by the market volatility of the S&P 500 for the 30 companies in the DJIA in the two-month period of December 2001 and January 2002. On average, the maximum divided by its minimum was 2.51 for the correlation, while it was only 1.28 for the ratio of volatilities. It is clear that the volatility of the betas is mainly a story of volatility of the correlations. Table 5 shows that, on average, the market price of the shares and the S&P 500 moved in the same direction (both increased or both decreased) only 58% of the months and 48.7% of the days in the 5-year period 1/1/1997-31/12/2001).

Table 5. Some statistics of the correlation and of the company volatility divided by the market volatility of the S&P500 for the 30 companies in the DJIA in the two-month period of December 2001 and January 2002.

On average, the maximum divided by its minimum was 2.51 for the correlation, while it was only 1.28 for the ratio of volatilities. Correlations and volatilities calculated using 60 monthly data.

$Beld_i = CO$	$deta_i = correlation (Return; Return (Market)) X (Con$				iipa	ny volatilit	yi / warket	voiatility)	
	Correlation				Company	volatility	/ Market vo	olatility	
	Average	Max	min	Max/min		Average	Max	min	Max/min
3M Co.	0.38	0.50	0.23	2.12		1.51	1.69	1.39	1.22
Alcoa	0.37	0.53	0.21	2.55		2.13	2.47	1.91	1.29
American Express	0.78	0.83	0.70	1.18		1.88	2.17	1.71	1.27
AT&T	0.34	0.44	0.14	3.11		2.22	2.57	1.98	1.29
Boeing	0.45	0.54	0.27	1.97		2.13	2.40	1.92	1.25
Caterpillar	0.40	0.52	0.31	1.65		1.93	2.15	1.74	1.24
Citigroup	0.80	0.85	0.71	1.19		2.07	2.32	1.89	1.23
Coca Cola	0.46	0.61	0.32	1.92		1.73	1.95	1.46	1.33
Du Pont	0.43	0.55	0.31	1.78		1.75	1.97	1.56	1.26
Eastman Kodak	0.28	0.35	0.18	1.97		1.98	2.25	1.67	1.35
Exxon Mobil	0.40	0.58	0.24	2.39		1.06	1.18	0.90	1.31
General Electric	0.79	0.83	0.74	1.11		1.59	1.72	1.46	1.18
Hewlett-Packard	0.53	0.68	0.39	1.76		2.55	2.90	2.23	1.30
Home Depot	0.62	0.71	0.50	1.43		2.01	2.28	1.82	1.25
Honeywell Intl.	0.50	0.59	0.42	1.40		2.44	2.76	2.29	1.20
IBM	0.56	0.66	0.36	1.83		2.11	2.30	1.89	1.22
Intel	0.57	0.64	0.50	1.28		2.72	3.09	2.43	1.27
Intl.Paper	0.38	0.48	0.29	1.67		1.90	2.15	1.65	1.30
Johnson & Johnson	0.37	0.53	0.25	2.14		1.29	1.50	1.10	1.37
J P Morgan Chase	0.72	0.77	0.66	1.18		2.21	2.50	1.89	1.32
McDonalds	0.41	0.49	0.30	1.64		1.60	1.85	1.50	1.23
Merck	0.36	0.55	0.14	4.02		1.67	1.89	1.42	1.33
Microsoft	0.60	0.66	0.52	1.26		2.55	3.10	2.24	1.38
Philip Morris	0.07	0.25	-0.07	3.67		1.96	2.20	1.66	1.32
SBC Comm.	0.30	0.47	0.15	3.16		1.62	1.83	1.36	1.35
United Technologies	0.64	0.72	0.56	1.28		1.89	2.12	1.70	1.25
Wal Mart Stores	0.56	0.67	0.46	1.45		1.78	2.06	1.62	1.27
General Motors	0.52	0.58	0.44	1.31		1.95	2.16	1.78	1.22
Procter & Gamble	0.24	0.45	-0.02	20.37		1.83	2.12	1.58	1.34
Walt Disney	0.55	0.70	0.43	1.61		1.90	2.07	1.69	1.22
Average	0.48	0.59	0.35	2.51		1.93	2.19	1.72	1.28
Max	0.80	0.85	0.74	20.37		2.72	3.10	2.43	1.38
Min	0.07	0.25	-0.07	1.11		1.06	1.18	0.90	1.18

Beta: = correlation (Return: : Return (Market)) x (Company volatility; / Market volatility)

2. Implications for making beta-ranked portfolios

We ordered the 3,813 companies by decreasing betas on December 1, 2001 and constructed 20 portfolios. Portfolio 1 had the companies with the highest betas and portfolio 20 had the companies with the lowest betas. Then we calculated the beta of the portfolios (weighted by market capitalization) each day of the following two months. **Table 7** shows that in the following two months, 300 portfolios were misallocated (i.e. on 26 days, portfolio 5 had lower beta than portfolio 6). On 53 days (out of 62 days) there were portfolios misallocated.

Having ordered the 3,813 companies by decreasing betas on December 1, 2001, we constructed 3,613 portfolios of 200 shares each following a moving window. We also calculated the beta of those 3,613 portfolios on December 15, 2001. **Figure 4** shows the results. Betas of low beta portfolios increased, and betas of high beta portfolios decreased from December 1 to December 15. Figure 5 shows the difference of the two betas (December 1 and December 15) for each portfolio. Figure 5 also shows the difference of the betas on December 15 between each portfolio (N) and the portfolio that had the immediate lower beta (N-1) in December 1. On December 15, this difference was negative in 1,520 cases.

We also formed portfolios in the Fama and French (1992) way on December 1 and on December 15, 2001. **Table 8** shows that on average 71.3% of the companies changed from one portfolio on December 1 to another on December 15.

	(1111777-51/12/2001)										
			All comp	panies	30 compar	nies DJIA					
Percentage range			Monthly data	Daily data	Monthly data	Daily data					
0	-	10%	0	0 10							
10%	-	20%	4	32							
20%	-	30%	7	126							
30%	-	40%	23	598							
40%	-	50%	404	1,138							
50%	-	60%	2,037	1,406	2						
60%	-	70%	1,227	474	16	24					
70%	-	80%	107	29	11	6					
80%	-	90%	4	0	1						
90%	-	100%	0	0							
Number	of co	ompanies	3,812	3,812	30	30					
A	vera	ge	58.0%	48.7%	68.3%	65.9%					
Median			58.1%	50.0%	66.9%	64.5%					

Table 6. Percentage days or months that the share price and the S&P 500 move in the same direction (1/1/1997-31/12/2001)

Table 7. Twenty portfolios ranked by decreasing beta on December 1, 2001 Number of misallocated portfolios on the 62 days of the following two months.

Date	misallocated
	portfolios
December: 10 and 16.	13
December: 12, 13, 14 and 15. January: 12, 13, 15 and 16.	11
January: 14.	10
December: 17 and 19.	9
December: 8 and 21. January: 8, 10, 17 and 19.	8
December: 18. January: 18, 20 and 21.	6
December: 5, 6, 7, 9, 11, 20 and 22. January: 5, 6, 7, 9, 11 and 22.	4
December: 3, 4, 23, 24, 27, 28, 29, 30 and 31. January: 3, 4, 23, 24, 27, 28, 29 and	2
30.	
9 days	0

Figure 4. 3,613 portfolios of 200 shares ranked by decreasing beta on December 1, 2001. Beta of the portfolios in December 1 (straight line) and beta of the same portfolios on December 15.



Figure 5. 3,613 portfolios of 200 shares ranked by decreasing beta on December 1, 2001. Difference of the beta of each portfolio in December 15 minus the beta of the same portfolio in December 1. The chart also shows the difference of the betas on December 15 between each portfolio and the portfolio that had the inmediate lower beta on December 1.



Table 8. Percentage of the companies in each portfolio formed on December 1 that change portfolio if portfolios formed on December 15, 2001. Change in the betas of each portfolio from December 1 to December 15, 2001. Portfolios formed according to Fama and French (1992).

Portfolios are formed on December 1 and on December 15, 2001. The breakpoints for the size (log of Market Value of Equity, ME, in million \$) are determined using all NYSE stocks (1,462) in our sample. All NYSE and Nasdaq stocks are allocated to the 10 size portfolios using the NYSE breakpoints. Then, each size decile is subdivided into 10 ß portfolios using the betas of individual stocks, estimated with 5 years of monthly returns ending in December 1, 2001 in the first case, and in December 15, 2001 in the second.

The betas of the portfolios are estimated with 5 years of monthly returns with respect to the S&P 500.

	All	Low-ß	ß-2	ß-3	ß-4	ß-5	ß-6	ß-7	ß-8	ß-9	High-ß
Panel A. Per	rcentage of t	he companie	es in eacl	h portfolio	formed o	n Decem	ber 1 that	change p	portfolio if	portfolios	s formed
	0			on E	December	15.		0.			
All	71.3%	52.8%	71.9%	78.2%	82.5%	79.3%	80.6%	82.8%	76.1%	71.1%	39.5%
Small-ME	74.7%	61.5%	77.8%	76.9%	83.8%	86.3%	79.5%	82.9%	77.8%	76.9%	45.2%
ME-2	73.3%	62.5%	73.4%	79.7%	87.5%	78.1%	78.1%	78.1%	81.3%	65.6%	48.4%
ME-3	73.1%	52.6%	76.3%	76.3%	81.6%	81.6%	86.8%	89.5%	84.2%	65.8%	40.9%
ME-4	72.9%	62.5%	75.0%	87.5%	71.9%	68.8%	81.3%	93.8%	75.0%	68.8%	45.5%
ME-5	66.6%	51.7%	65.5%	72.4%	82.8%	72.4%	75.9%	86.2%	62.1%	69.0%	31.3%
ME-6	68.4%	47.6%	61.9%	81.0%	85.7%	90.5%	76.2%	66.7%	81.0%	66.7%	30.4%
ME-7	62.6%	19.0%	42.9%	61.9%	81.0%	71.4%	90.5%	85.7%	71.4%	76.2%	27.3%
ME-8	65.4%	31.6%	73.7%	84.2%	78.9%	63.2%	78.9%	78.9%	68.4%	68.4%	30.0%
ME-9	67.5%	26.3%	63.2%	84.2%	78.9%	84.2%	84.2%	84.2%	68.4%	84.2%	26.1%
Large-ME	66.1%	41.2%	76.5%	82.4%	82.4%	70.6%	82.4%	76.5%	70.6%	58.8%	22.2%
						-				_	
	Panel B	. Portfolio we	eighted B	eta Dece	mber, 1 -	Portfolio	weighted	Beta Dec	ember, 18)	
	All	Low-ß	ß-2	ß-3	ß-4	ß-5	ß-6	ß-7	ß-8	ß-9	High-ß
All		-0.34	-0.22	-0.17	-0.20	-0.09	-0.09	-0.03	0.04	0.04	0.27
Small-ME	-0.09	-0.42	-0.22	-0.17	-0.17	-0.10	-0.13	0.04	0.06	-0.03	0.20
ME-2	-0.11	-0.36	-0.24	-0.21	-0.20	-0.07	-0.14	-0.14	-0.05	0.06	0.26
ME-3	-0.11	-0.29	-0.29	-0.23	-0.26	-0.09	-0.13	-0.18	0.07	0.04	0.19
ME-4	-0.11	-0.48	-0.25	-0.22	-0.17	-0.12	-0.09	-0.03	-0.03	-0.12	0.42
ME-5	-0.06	-0.35	-0.26	-0.17	-0.26	0.01	-0.04	-0.04	0.05	0.15	0.30
ME-6	-0.06	-0.27	-0.16	-0.16	-0.12	-0.17	-0.02	0.00	0.03	0.00	0.25
ME-7	-0.03	-0.20	-0.17	-0.11	-0.24	-0.08	-0.01	-0.01	0.10	0.24	0.19
ME-8	0.01	-0.13	-0.17	-0.06	-0.18	-0.09	-0.03	0.08	0.21	0.09	0.41
ME-9	-0.01	-0.22	-0.14	-0.16	-0.15	-0.16	-0.08	0.12	0.05	0.13	0.42
Large-ME	-0.01	-0.19	-0.18	-0.06	-0.24	-0.07	0.06	-0.03	0.08	0.22	0.34

3. Historical betas depend very much on which index we use to calculate them

Table 9 presents the historical relative betas of the 30 companies in the Dow Jones Industrial Average Index. Relative betas are calculated by dividing the beta with respect to an index on a given day by the beta with respect to another index on the same day. For the 2-month period 1/12/01-31/1/02, the table contains the maximum, the minimum, the average, and maximum divided by the minimum. It may be seen that, on average, the beta with respect to the S&P 500 was smaller than the beta with respect to the DJIA and higher than the beta with respect to the W5000. Table 9 permits to conclude that relative betas also change dramatically.

Table 9. Historical relative betas of the 30 companies in the Dow Jones Industrial Average Index

Relative betas are calculated by dividing the beta with respect to one index on a given day by the beta with respect to another index on the same day. For example, the relative beta "Beta S&P 500 / Beta DJ IND" is calculated by dividing the beta with respect to the S&P 500 on a given day by the beta with respect to the Dow Jones Industrial Average (DJIA) index on the same day.

calculated beta with respect to the DJIA on December 18, 2001

The table contains the maximum, the minimum, the average, and the maximum divided by the minimum of the 62 relative betas calculated in the period 1/12/01-31/1/02

	Dotas	Juiculation	5 u in the		12/01 0	11 11 02			
	Beta S&P 500 / Beta DJ IND					Beta S&P 500 / Beta W 5000			
	Max	min a	average	max/min		Max	min	average	max/min
3M Co.	0.76	0.51	0.65	1.51		2.02	1.00	1.20	2.02
Alcoa	0.86	0.46	0.66	1.86		1.22	0.94	1.03	1.30
American Express	1.10	0.89	0.99	1.23		1.16	1.04	1.10	1.11
AT&T	2.57	1.19	1.66	2.15		1.09	0.92	1.00	1.19
Boeing	0.84	0.55	0.74	1.51		1.11	0.99	1.05	1.13
Caterpillar	0.78	0.48	0.66	1.62		1.37	1.10	1.21	1.24
Citigroup	1.18	0.96	1.05	1.24		1.16	1.02	1.09	1.13
Coca Cola	1.02	0.80	0.90	1.26		1.30	1.07	1.17	1.21
Du Pont	0.81	0.54	0.70	1.51		1.75	1.09	1.26	1.60
Eastman Kodak	0.90	0.56	0.75	1.60		1.24	0.92	1.10	1.35
Exxon Mobil	0.91	0.69	0.80	1.32		1.38	0.97	1.14	1.43
General Electric	1.23	1.00	1.10	1.23		1.21	1.07	1.12	1.13
Hewlett-Packard	1.33	1.06	1.19	1.26		1.01	0.85	0.93	1.19
Home Depot	1.30	1.00	1.15	1.30		1.18	0.99	1.08	1.19
Honeywell Intl.	0.87	0.68	0.80	1.29		1.31	1.09	1.18	1.20
IBM	1.10	0.80	0.98	1.38		1.10	0.99	1.05	1.12
Intel	1.38	1.06	1.24	1.30		1.04	0.95	1.00	1.09
Intl.Paper	0.80	0.49	0.66	1.62		1.41	1.10	1.22	1.29
Johnson & Johnson	1.12	0.60	0.90	1.86		1.90	1.22	1.40	1.56
J P Morgan Chase	1.24	1.02	1.11	1.21		1.08	0.96	1.04	1.12
McDonalds	1.12	0.73	0.93	1.53		1.81	1.19	1.36	1.51
Merck	1.29	0.70	1.06	1.84		2.67	1.27	1.54	2.11
Microsoft	1.55	1.02	1.34	1.52		1.09	0.99	1.04	1.10
Philip Morris	12.1	-28.1	-0.07	0.43		27.1	-2.86	1.80	9.47
SBC Comm.	1.94	0.98	1.46	1.97		1.74	1.07	1.26	1.62
United Technologies	0.92	0.77	0.86	1.19		1.23	1.07	1.14	1.15
Wal Mart Stores	1.25	0.96	1.10	1.31		1.31	1.11	1.22	1.18
General Motors	1.07	0.85	0.98	1.26		1.07	0.95	1.01	1.13
Procter & Gamble	1.00	-0.16	0.63	6.22		3.09	0.49	1.49	6.34
Walt Disney	1.21	0.90	1.05	1.33		1.14	0.96	1.06	1.18
Average	1.52	0.80	0.97	1.59		2.28	1.01	1.18	1.75

4. We cannot say that the beta of a company is smaller or bigger than the beta of another

Table 10 presents the beta ranking of the 3,813 companies in our sample in the month of December 31, 2001. Each day, companies are ranked from 1 (the company with the lowest beta on that day) to 3,813 (the company with the highest beta on that day). Betas are calculated each day with respect to the S&P 500 using 5 years of monthly data. It may be seen that the average change in ranking for all 3,813 companies in December 2001 is 1,542 ranking positions.

The average beta ranking change was: 233 positions from one day to the next; 479 positions from one day to the next week; and 564 positions over a two-week period.

Table 10. Change in beta ranking order in the month of December, 2001. Statistics of the difference Maximum beta ranking - minimum beta ranking.

Historical betas of 3,813 companies calculated every day during the month of December 2001 with respect to the S&P 500 using 5 years of monthly data. Each day, companies are assigned a beta ranking from 1 (the company with the minimum beta) to 3,813 (the company with the maximum beta). Then, we calculate for each company the difference between the Maximum beta ranking and the minimum beta ranking.

Maximum ranking - minimum ranking	Full sample	S&P 500	DJIA 30
MAX	3,760	2,592	2,041
Min	15	74	467
Average	1,542	1,154	1,001
Median	1,391	1,126	908
Number of companies	3,813	450	30

5. High-risk companies very often have smaller historical betas than low-risk companies

Table 11 reports the calculated betas as of December 31, 2001 of the 30 companies in the Dow Jones Industrial Average Index. Companies are sorted by ascending beta with respect to the S&P 500. According to the S&P 500 betas, Philip Morris is the company with lowest cost of equity, much smaller than GE or Wall Mart. If we assume that the risk-free rate is 5%, that the market risk premium is 4.5%, and that historical betas are a good proxy for expected betas, then the cost of equity of Philip Morris, GE and Wall Mart is 6.0%, 10.2% and 9.1%, respectively. We do not think that this makes much economic sense.

Table 11. Calculated betas as of December	31, 2001 of the	30 companies in the	Dow Jones Industrial	Average Index.
Betas are calculated each day using 5 years	of monthly data.	Companies are sorted	I by ascending beta with	n respect to the

	31/12/2001	Beta S&P 500	Beta DJ IND	Beta W 5000
PHILIP MORRIS	MO	0.232	0.378	0.149
PROCTER & GAMBLE	PG	0.281	0.460	0.244
EXXON MOBIL	XOM	0.358	0.480	0.329
SBC COMMUNICATIONS	SBC	0.460	0.367	0.322
MERCK	MRK	0.483	0.528	0.299
3M CO.	MMM	0.488	0.778	0.464
JOHNSON & JOHNSON	JNJ	0.488	0.578	0.363
EASTMAN KODAK	EK	0.590	0.791	0.522
COCA COLA	KO	0.654	0.775	0.530
MCDONALDS	MCD	0.677	0.769	0.500
CATERPILLAR	CAT	0.731	1.104	0.633
DU PONT	DD	0.771	1.019	0.700
BOEING	BA	0.807	1.178	0.744
WAL MART STORES	WMT	0.917	0.820	0.749
WALT DISNEY	DIS	0.928	1.022	0.844
AT&T	Т	0.959	0.542	0.942
INTL.PAPER	IP	1.011	1.323	0.877
GENERAL MOTORS	GM	1.129	1.167	1.087
HOME DEPOT	HD	1.130	0.891	1.143
GENERAL ELECTRIC	GE	1.163	0.994	1.063
HONEYWELL INTL.	HON	1.176	1.564	1.016
ALCOA	AA	1.219	1.487	1.209
IBM	IBM	1.234	1.196	1.171
AMERICAN EXPRESS	AXP	1.245	1.374	1.120
UNITED TECHNOLOGIES	UTX	1.323	1.488	1.207
CITIGROUP	С	1.459	1.525	1.320
HEWLETT-PACKARD	HPQ	1.489	1.270	1.639
J P MORGAN CHASE & CO.	JPM	1.518	1.431	1.458
INTEL	@INTC	1.696	1.434	1.632
MICROSOFT	@MSFT	1.823	1.379	1.716
average		0.948	1.004	0.866

S&P 500.

6. Weak correlation between beta and realized return

 Table 12 shows the small correlation between the betas and the realized returns of portfolios of 200 companies sorted by realized return.

Table 12. Portfolios of 200 companies sorted by realized return. Correlation of the portfolio return with the beta calculated on December 1, 2001

	CO	CORRELATION REALIZED RETURN - BETA.												
	1996-2001	1997	1998	1999	2000	2001								
Slope	-0.07	0.23	0.64	1.68	-1.00	-0.26								
standard error slope	0.076	0.032	0.038	0.013	0.009	0.023								
\mathbb{R}^2	0.0002	0.0145	0.0728	0.8155	0.7667	0.0347								
F statistic	0.8	53.2	282.6	15917.3	11835.9	129.4								
intercept	0.54	0.02	-0.61	-1.25	1.10	0.43								
Returns:														
S&P 500	66.2%	33.4%	28.6%	21.0%	-9.1%	-11.9%								
DOW JONES IND.	68.8%	24.9%	18.1%	26.7%	-4.5%	-5.4%								
WILSHIRE 5000	48.8%	29.2%	21.7%	22.0%	-11.8%	-12.1%								

7. About the recommendation of using Industry betas

Some authors recommend using industry betas, instead of company betas. For example, Copeland, Koller and Murrin (2000) recommend "checking several reliable sources because beta estimates vary considerably... If the betas from several sources vary by more than 0.2 or the beta for a company is more than 0.3 from the industry average, consider using the industry average. An industry average beta is typically more stable and reliable than an individual company beta because measurement errors tend to cancel out". But about the CAPM, they conclude (see their page 225), "It takes a better theory to kill an existing theory, and we have not seen the better theory yet. Therefore, we continue to use the CAPM, being wary of all the problems with estimating it." We rather think that a rejection is enough to beat a model.

We have shown that industry betas are quite unstable.

Figure 6. Betas of the banking industry. 409 companies. Market capitalization: \$1,150 bn.







Figure 6 shows the calculated betas of the banking industry in the period December 1, 2001 - January 31, 2002. It shows the evolution of the industry betas (both weighted and unweighted), the maximum beta and the minimum beta of the 409 companies in the industry. The difference between the weighted and unweighted betas is remarkable.

Figure 7 shows the average beta (in the period December 1, 2001 – January 31, 2002), the maximum beta and the minimum beta for each of the 409 companies in the banking industry.

Figure 8 shows the relationship between the average beta (in the period December 1, 2001 – January 31, 2002) and the market capitalization of the 409 companies in the banking industry. On average, bigger companies had higher betas.

Figure 9 shows the average and the dispersion of the historical betas of the 30 banks with the highest market capitalization and compares them with the average and the dispersion of the industry weighted and unweighted betas.



Figure 8. Betas of the banking industry. Average beta and Market capitalization of the 409 banks

Figure 9. Betas of the banking industry. Dispersion of the betas of the 30 banks with the highest market cap.



8. Historical betas and the market-to-book ratio.

Figure 10 shows the relationship between the market-to-book ratio and the historical beta of portfolios of 200 companies ranked by MV/BV (market-to-book ratio). The figure plots the 200 companies' rolling average beta on December 1 and on December 15. On average, higher MV/BV (market-to-book ratio) companies had higher beta.

Figure 11 shows the relationship between the market-to-book ratio and the market capitalization of portfolios of 200 companies ranked by market capitalization. The figure plots the 200 companies rolling average MV/BV (market-to-book ratio) on December 1 and on December 15. On average, bigger companies had higher MV/BV (market-to-book ratio).

Figure 12 shows the relationship between the market capitalization and the calculated beta of portfolios of 200 companies ranked by market capitalization on December 1. Small caps had low betas.









Figure 12. Relationship between market capitalization and calculated beta. Portfolios of 200 companies ranked by market capitalization on December 1



9. Conclusion

Pablo Fernandez

We have shown that, in general, it is an enormous error to use the historical beta as a proxy for the expected beta. First, because it is almost impossible to calculate a meaningful beta because historical betas change dramatically from one day to the next; second, because very often we cannot say with a relevant statistical confidence that the beta of one company is smaller or bigger than the beta of another; third, because historical betas do not make much sense in many cases: high-risk companies very often have smaller historical betas than low-risk companies; fourth, because historical betas depend very much on which index we use to calculate them.

We calculate betas of 3,813 companies using 60 monthly returns each day of December 2001 and January 2002. We report that the maximum beta of a company was, on average, 15.7 times its minimum beta. The median of the maximum beta divided by the minimum beta was 3.07. The median of the percentage daily change (in absolute value) of the betas was 20%, and the median of the percentage (in absolute value) of the betas was 43%.

Industry betas are also unstable. The median (average) of the percentage daily change (in absolute value) of the industry betas was 7% (16%), and the median (average) of the percentage (in absolute value) of the industry betas was 15% (38%). On average, the maximum beta of an industry was 2.7 times its minimum beta in December 2001 and January 2002.

This dispersion of the calculated betas also has important implications for the instability of beta-ranked portfolios.

Appendix 1. Short literature review about the CAPM

Sharpe (1964) and Lintner (1965) demonstrate that, in equilibrium, a financial asset's return must be positively linearly related to its beta (B), a measure of systematic risk or co-movement with the market portfolio return:

 $E(R_i) = a_1 + a_2 E(\beta_i)$, for all assets i,

(1)

where E (R_i) is the expected return on asset i, E (β_i) is asset i's expected market beta, a_1 is the expected return on a "zero-beta" portfolio, and a_2 is the market risk premium.

The CAPM of Sharpe (1964), Lintner (1965) and Mossin (1966) is predicated on the assumption of a positive systematic risk-return tradeoff and asserts that the expected return for any security is a positive function of three variables: expected beta, expected market return, and the risk-free rate. The basic assumptions of the CAPM are:

1. Investors have homogeneous expectations about asset returns that have a joint normal distribution;

- 2. Investors are risk-averse individuals who maximize the expected utility of their end-of-period wealth;
- 3. Markets are frictionless and information is costless and simultaneously available to all investors; there are no imperfections such as taxes, regulations, or restrictions on short selling;
- 4. There exists a risk-free asset such that investors may borrow and lend unlimited amounts at the risk-free rate.

However, subsequent work by (among many others) Basu (1977), Banz (1981), Reinganum (1981), Litzenberger and Ramaswamy (1979), Keim (1983, 1985)² and Fama and French (1992) suggests that either:

- 1. expected returns are determined not only by the beta and the expected market risk premium but also by non-risk characteristics such as book-to-market ratio, firm size, price-earnings ratio and dividend vield. It implies that the CAPM is misspecified and requires the addition of factors other than beta to explain security returns, or
- 2. the historical beta has little (or nothing) to do with the expected beta. To put it another way: the problems of measuring beta are systematically related to variables such as firm size and book-to-market ratio. And also the historical market risk premium has little (or nothing) to do with the expected market risk premium, or
- 3. the heterogeneity of expectations³ in cross-section returns, volatilities and covariances, and market returns is the reason why it makes no sense to talk about an aggregate market CAPM, although at the individual level expected CAPM does

² Basu (1977) found that low price/earnings portfolios have higher returns than could be explained by the CAPM. Banz (1981) and Reinganum (1981) found that smaller firms tend to have high abnormal rates of return. Litzenberger and Ramaswamy (1979) found that the market requires higher rates of return on equities with high dividend yield. Keim (1983, 1985) reports the January effect, that is, seasonality in stock returns. Tinic and West (1984) reject the validity of the CAPM based on intertemporal inconsistencies due to the January effect.

³ Lintner (1969) argued that the existence of heterogeneous expectations does not critically alter the CAPM in some simplified scenarios. In some cases, expected returns are expressed as complex weighted averages of investors expectations. But if investors have heterogeneous expectations of expected prices and covariance matrix, the market portfolio is not necessarily efficient and this makes the CAPM non-testable. Lintner (1969) says "in the (undoubtedly more realistic) case with different assessments of covariance matrices, the market's assessment of the expected ending price for any security

Pablo Fernandez IESE Business School, University of Navarra

work. It means that while individuals are well characterized by CAPM, and each individual uses an expected beta, an expected market risk premium, and an expected cash flow stream to value each security, all individuals do not agree on these three magnitudes for each security. Consequently, it makes no sense to refer to a "market" expected beta for a security or to a "market" expected market risk premium (or to a "market" expected cash flow stream), for the simple reason that they do not exist.

We may find out an investor's expected IBM beta by asking him. However, it is impossible to determine the expected IBM beta for the market as a whole, because it does not exist. Even if we knew the expected market risk premiums and the expected IBM betas of the different investors who operated on the market, it would be meaningless to talk of an expected IBM beta for the market as a whole. The rationale for this is to be found in the aggregation theorems of microeconomics, which in actual fact are non-aggregation theorems. A model that works well individually for a number of people may not work for all of the people together⁴. For the CAPM, this means that although the CAPM may be a valid model for each investor, it is not valid for the market as a whole, because investors do not have the same return and risk expectations for all shares. The prices are a statement of expected cash flows discounted at a rate that includes the expected market risk premium and the expected beta. Different investors have different cash flow expectations and different future risk expectations (different expected market risk premium and different expected beta). One could only talk of a market risk premium if all investors had the same expectations.

The problem with the expected beta is that investors do not have homogeneous expectations. If they did, it *would* make sense to talk of a market risk premium and of an IBM beta common to all investors because all investors would hold the market portfolio. However, expectations are not homogeneous.

САРМ	Real world
Homogeneous expectations All investors have equal expectations about asset returns that have a joint normal distribution	Heterogeneous expectations. All investors DO NOT have equal expectations about asset returns. Asset returns DO NOT have a joint normal distribution
All investors use the same beta for each share	Investors use different betas (required betas) for a share
All investors hold the market portfolio	Investors hold different portfolios
All investors have the same expected market risk premium	Investors have different expected market risk premia and use different required market risk premia
The market risk premium is the difference between the <i>expected</i> return on the market portfolio and the risk-free rate	The market risk premium is NOT the difference between the <i>expected</i> return on the market portfolio and the risk-free rate

Measurement errors and problems

Original tests of the CAPM focused on whether the intercept in a cross-sectional regression was higher or lower than the risk-free rate, and whether stock individual variance entered into cross-sectional regressions.

Scholes and Williams (1977) found that with nonsyncronous trading of securities, ordinary least squares estimators of beta coefficients using daily data are both biased and inconsistent.

Roll (1977) concludes that the only legitimate test of the CAPM is whether or not the market portfolio (which includes all assets) is mean-variance efficient. The Roll critique does not imply that the CAPM is an invalid theory. However, it does mean that tests of the CAPM must be interpreted with great caution.

Roll (1981) suggests that infrequent trading of shares of small firms may explain much of the measurement error in estimating their betas.

Constantinides (1982) points out that with consumer heterogeneity "in the intertemporal extension of the Sharpe-Lintner CAPM, an asset's risk premium is determined not only by its covariance with the market return, but also by its covariance with the m-1 state variables" (m is the number of heterogeneous consumers). He also points out that the assumption of complete markets is needed for demand aggregation. But markets are not complete.

Lakonishok and Shapiro (1984, 1986) find an insignificant relationship between beta and returns and a significant relationship between market capitalization and returns

Shanken (1992) presents an integrated econometric view of maximum-likelihood methods and two-pass approaches to estimating historical betas.

depends on every investor's assessment of the expected ending price for every security and every element in the investor's assessment of his NxN covariance matrix (N is the number of securities), as well as the risk tolerance of every investor."

⁴ As Mas-Colell el al. (1995, page 120) say: *"It is not true that whenever aggregate demand can be generated by a representative consumer, this representative consumer's preferences have normative contents. It may even be the case that a positive representative consumer exists but that there is no social welfare function that leads to a normative representative consumer."*

Roll and Ross (1994) attribute the observed lack of a systematic relation between risk and return to the possible mean-variance inefficiency of the market portfolio proxies.

Shalit and Yitzhaki (2002) argue that the Ordinary Least-Squares regression estimator is inappropriate for estimating betas. They suggest alternative estimators for beta that are robust with respect to extreme fluctuations in the market return. Using CRSP daily data from 1984 to 1993, they eliminate the highest four and the lowest four market returns and show that the betas of the 75% of the firms change by more than one standard error. For example, the beta of General Electric changes from 1.16 to 1.23 and the beta of Coca-Cola changes from 1.22 to 1.18.

Avramov (2002) uses Bayesian model averaging to analyze the sample evidence on return predictability in the presence of model uncertainty. The analysis reveals in-sample and out-of-sample predictability, and shows that the out-of-sample performance of the Bayesian approach is superior to that of model selection criteria. She finds that market premia are robust predictors. Moreover, small-cap value stocks appear more predictable than large-cap growth stocks. She also investigates the implications of model uncertainty from investment management perspectives. She shows that model uncertainty is more important than estimation risk, and investors who discard model uncertainty face large utility losses.

Zhang, Kogan, and Gomes (2001) reconcile the ability of non-risk characteristics such as firm size and book-tomarket to predict returns within a dynamic pricing paradigm. Firm characteristics can appear to predict stock returns because they may be correlated with the true conditional factor loadings, thereby motivating the scaling of betas by firm specific variables. They claim that "size and book-to-market play separate roles in describing the cross-section of returns. These firm characteristics appear to predict stock returns because they are correlated with the true conditional market beta of returns."

Avramov and Chordia (2001) test whether the Zhang, Kogan, and Gomes (2001) scaling procedure improves the performance of the theoretically motivated CAPM and consumption CAPM. The evidence shows that equity characteristics often enter beta significantly. However, "characteristic scaled factor models" do not outperform their unscaled counterparts.

The poor performance of the CAPM has inspired multiple portfolio based factors.

The article that dealt the hardest blow to the CAPM was that published by Fama and French (1992). This article showed that in the period 1963-1990, the correlation between stocks' returns and their betas was very small, while the correlation with the companies' size and their price/book value ratio was greater. They concluded "our tests do not support the most basic prediction of the Sharpe-Lintner-Black Capital Asset Pricing Model that average stock returns are positively related to market betas". The authors divided the shares into portfolios and found that the cross-sectional variation in expected returns may be captured within a three-factor model, the factors being:

- the return on the market portfolio in excess of the risk-free rate

- a zero net investment portfolio that is long in high book-to-market stocks and short in low book-to-market stocks.

- a zero net investment portfolio that is long in small firm stocks and short in large firm stocks.
 - Table A.1 shows the article's main findings.

Size		Annual	Beta		Annual			Annual
of the	Average	average	of the	Average	average	Price /	Average	average
companies	beta	return	companies	beta	return	book value	beta	return
1 (biggest)	0.93	10.7%	1 (high)	1.68	15.1%	1 (high)	1.35	5.9%
2	1.02	11.4%	2	1.52	16.0%	2	1.32	10.4%
3	1.08	13.2%	3	1.41	14.8%	3	1.30	11.6%
4	1.16	12.8%	4	1.32	14.8%	4	1.28	12.5%
5	1.22	14.0%	5	1.26	15.6%	5	1.27	14.0%
6	1.24	15.5%	6	1.19	15.6%	6	1.27	15.6%
7	1.33	15.0%	7	1.13	15.7%	7	1.27	17.3%
8	1.34	14.9%	8	1.04	15.1%	8	1.27	18.0%
9	1.39	15.5%	9	0.92	15.8%	9	1.29	19.1%
10 (smallest)	1.44	18.2%	10 (low)	0.80	14.4%	10 (low)	1.34	22.6%

Table A.1. Main findings of Fama and French's article (1992)

However, Griffin (2002) concludes that there are no benefits to extending the Fama and French three-factor model to a global context. Country-specific three-factor models are more useful in explaining stock returns than are world and international versions.

Lakonishok, Shleifer and Vishny (1994) argue that the size and book-to-market effects are due to investor overreaction rather than compensation for risk bearing. According to them, investors systematically overreact to corporate news, unrealistically extrapolating high or low growth into the future. This leads to underpricing of "value" (small capitalization, high book-to-market stocks) and overpricing of "growth" (large capitalization, low book-to-market stocks).

Kothary, Shanken and Sloan (1995) point out that using historical betas estimated from annual rather than monthly returns produces a stronger relation between average return and historical beta. They also claim that the relation between book-to-market equity and average return observed by Fama and French (1992) and others is seriously exaggerated by survivor bias in the COMPUSTAT sample used. They also claim that the Fama and French statistical tests were of such low power that they could not reject a beta-related risk premium of 6% over the post-1940 period. Their most important

conclusion, however, is that "our examination of the cross-section of expected returns reveals economically and statistically significant compensation (about 6 to 9% per annum) for beta risk."

Pettengill, Sundaram and Mathur (1995) find "consistent and highly significant relationship between beta and crosssectional portfolio returns". They insist that "the positive relationship between returns and beta predicted by CAPM is based on expected rather than realized returns. In periods where excess market returns are negative, an inverse relationship between beta and portfolio returns should exist". They test the following equation:

 $R_{it} = \gamma_{0t} + \gamma_{1t} \ \delta \ \beta_i + \gamma_{2t} \ (1 - \delta) \ \beta_i + \epsilon_t$

where $\delta = 1$ if $(\mathbf{R}_{Mt} - \mathbf{R}_{Ft}) > 0$ (when market excess returns are positive) and $\delta = 0$ if $(\mathbf{R}_{Mt} - \mathbf{R}_{Ft}) < 0$ (when market excess returns are negative). γ_{1t} is estimated in periods with positive market excess returns and γ_{2t} is estimated in periods with negative market excess returns.

They use 660 monthly CRSP returns from 1936 to 1990. They estimate the betas using a five-year period and the CRSP equally-weighted index as a proxy for the market index. Based on the relative rankings of the estimated betas, securities are divided into 20 portfolios. They find a mean value of 0.0336 (t-statistic = 12.61) for γ_1 and -0.0337 (t-statistic = -13.82) for γ_2 . They say "as expected, high beta portfolios incur lower returns during down markets (280 months) than low beta portfolios... and receive a positive risk premium during up markets (380 months)." They also remark that their results are very similar to those of Lakonishok and Shapiro (1984), who found slope coefficients of $\gamma_1 = -0.0333$ and $\gamma_2 = -0.0354$.

Elsas, El-Shaer and Theissen (2000) follow the Pettengill, Sundaram and Mathur (1995) methodology for the German market and find a positive and statistically significant relation between beta and return in our sample period 1960-1995 as well as in all subperiods we analyze. They claim, *"Our empirical results provide a justification for the use of betas estimated from historical return data by portfolio managers."*

Carelli *et al.* (2014)⁵ calculate betas of 1,385 US companies on March 31, 2014: "147 betas for each company using monthly, weekly and daily returns and using different intervals: from 1 to 5 years. The median of the difference [maximum beta - minimum beta] was 1.03. Ranking the companies according to their betas, we find that the average of the [maximum ranking – minimum ranking] for the 1,385 companies is 786."



Almost all of the papers about CAPM published in journals in the last 48 years relay in <u>one</u> calculated beta <u>per</u> <u>date</u>. Carelli *et al.* (2014) show that for a single date, calculated betas have an average range of 1.03. It is easy to contrast that Carelli's numbers are correct: they imply that most papers that use calculated betas are irrelevant. How is it possible that so many very intelligent referees and editors have approved the publication of so many papers during so many years?

Fernandez (2015a) claims that "the CAPM is an absurd model because its assumptions and its predictions/conclusions have no basis in the real world". He also points out that "The CAPM is about expected return" and raises two questions: "If you find a formula for expected returns that works well in the real markets, would you publish it? Before or after becoming a billionaire?"

⁵ "Which is the right "*Market Beta*?", <u>http://ssrn.com/abstract=2509849</u>). The authors calculate betas of 1,385 US companies on March 31, 2014: 147 betas for each company.

Pablo Fernandez IESE Business School, University of Navarra

Appendix 2. Summary statistics of the historical betas of the 30 companies in the Dow Jones Industrial Average Index with respect to the S&P 500, the DJIA and the W 5000. Betas are calculated each day in the period 1/12/01-31/1/02 using 5 years of monthly data, i.e. on December 18, 2001, the beta is calculated by running a regression of the 60 monthly returns of the corresponding index. The returns of each month are calculated on the 18^{th} of the month. The table contains the maximum, the minimum, the average, maximum minus minimum divided by the average, and maximum divided by the minimum of the 62 betas calculated every day in the period 1/12/01-31/1/02.

[Beta S&	&P 500		Beta DJIA					Beta W 5000					Volatility				
1/12/01 -31/1/02	Max	min	average	(max-min)/	max/min	Max	min	average	(max-min)/	max/min	Max	min	average	(max-min)/	max/min	Max	min	average	(max-min)/	max/min
				average					average					average					average	
MMM	0.79	0.36	0.57	0.76	2.22	1.09	0.69	0.87	0.46	1.58	0.69	0.18	0.49	1.06	3.93	0.31	0.25	0.27	0.23	1.25
AA	1.22	0.45	0.78	0.99	2.73	1.49	0.96	1.18	0.45	1.55	1.21	0.44	0.77	1.00	2.72	0.42	0.36	0.38	0.17	1.19
AXP	1.76	1.24	1.46	0.36	1.42	1.75	1.33	1.48	0.28	1.32	1.62	1.11	1.33	0.38	1.46	0.39	0.29	0.34	0.29	1.34
T	1.02	0.32	0.77	0.91	3.16	0.75	0.15	0.48	1.24	4.94	1.03	0.35	0.77	0.87	2.92	0.45	0.36	0.40	0.24	1.26
BA	1.22	0.57	0.95	0.68	2.14	1.60	1.00	1.28	0.47	1.59	1.17	0.54	0.91	0.69	2.15	0.44	0.35	0.38	0.23	1.25
CAT	1.04	0.60	0.78	0.58	1.75	1.36	0.96	1.18	0.34	1.41	0.85	0.49	0.64	0.56	1.74	0.38	0.31	0.35	0.18	1.20
С	1.96	1.45	1.66	0.31	1.35	1.71	1.49	1.58	0.14	1.15	1.79	1.31	1.52	0.32	1.37	0.43	0.34	0.37	0.24	1.26
КО	1.11	0.55	0.80	0.70	2.02	1.18	0.67	0.88	0.58	1.76	0.97	0.44	0.68	0.78	2.21	0.34	0.27	0.31	0.22	1.25
DD	0.87	0.54	0.74	0.45	1.62	1.19	0.93	1.07	0.25	1.28	0.76	0.38	0.60	0.63	2.00	0.35	0.28	0.31	0.25	1.28
EK	0.70	0.37	0.56	0.60	1.92	0.93	0.61	0.75	0.42	1.51	0.70	0.31	0.51	0.77	2.27	0.40	0.31	0.35	0.24	1.28
XOM	0.67	0.24	0.43	1.00	2.81	0.79	0.34	0.54	0.85	2.35	0.61	0.22	0.38	1.05	2.82	0.22	0.17	0.19	0.28	1.31
GE	1.39	1.12	1.25	0.22	1.25	1.35	0.94	1.14	0.36	1.43	1.25	0.97	1.11	0.25	1.29	0.31	0.26	0.28	0.20	1.22
HPQ	1.69	0.94	1.34	0.56	1.80	1.36	0.86	1.13	0.44	1.58	1.74	0.98	1.43	0.53	1.78	0.50	0.41	0.46	0.19	1.21
HD	1.44	0.99	1.24	0.36	1.45	1.23	0.89	1.08	0.31	1.38	1.39	0.85	1.15	0.47	1.64	0.41	0.33	0.36	0.24	1.27
HON	1.42	1.03	1.22	0.32	1.39	1.74	1.26	1.52	0.32	1.39	1.28	0.79	1.04	0.47	1.62	0.48	0.41	0.44	0.16	1.17
IBM	1.36	0.76	1.18	0.51	1.79	1.45	0.92	1.19	0.45	1.58	1.33	0.73	1.12	0.53	1.82	0.44	0.34	0.38	0.25	1.27
@INTC	1.85	1.25	1.55	0.39	1.48	1.53	0.92	1.25	0.49	1.66	1.86	1.30	1.55	0.36	1.43	0.55	0.42	0.49	0.26	1.30
IP	1.02	0.49	0.73	0.73	2.09	1.33	0.93	1.10	0.37	1.43	0.88	0.40	0.60	0.81	2.23	0.38	0.30	0.34	0.23	1.25
JNJ	0.75	0.33	0.47	0.89	2.26	0.71	0.41	0.52	0.56	1.70	0.56	0.18	0.34	1.11	3.17	0.27	0.20	0.23	0.32	1.37
JPM	1.93	1.38	1.59	0.34	1.39	1.76	1.25	1.44	0.36	1.41	1.92	1.32	1.54	0.39	1.46	0.45	0.34	0.40	0.27	1.32
MCD	0.81	0.48	0.66	0.50	1.70	0.90	0.53	0.72	0.51	1.68	0.68	0.27	0.50	0.83	2.55	0.31	0.27	0.29	0.14	1.15
MRK	0.95	0.24	0.61	1.18	4.04	0.80	0.31	0.56	0.88	2.58	0.71	0.09	0.42	1.47	8.10	0.33	0.27	0.30	0.21	1.24
@MSFT	1.91	1.27	1.53	0.42	1.51	1.43	0.82	1.15	0.54	1.76	1.78	1.22	1.47	0.38	1.46	0.52	0.39	0.46	0.28	1.32
MO	0.42	-0.15	0.12	4.70		0.59	-0.01	0.31	1.93		0.31	-0.25	0.00	474.83		0.39	0.30	0.35	0.24	1.28
SBC	0.82	0.23	0.48	1.23	3.59	0.52	0.18	0.33	1.04	2.97	0.75	0.14	0.40	1.54	5.37	0.34	0.24	0.29	0.33	1.40
UTX	1.39	0.98	1.20	0.35	1.43	1.65	1.09	1.40	0.40	1.52	1.25	0.84	1.06	0.39	1.49	0.37	0.30	0.34	0.19	1.21
WMT	1.20	0.77	0.99	0.43	1.55	1.11	0.67	0.90	0.49	1.66	1.04	0.60	0.82	0.54	1.75	0.36	0.28	0.32	0.26	1.30
GM	1.20	0.78	1.01	0.41	1.52	1.22	0.88	1.03	0.33	1.38	1.18	0.75	1.00	0.43	1.58	0.38	0.33	0.35	0.14	1.15
PG	0.91	-0.04	0.45	2.10		1.07	0.24	0.66	1.26	4.49	0.61	-0.08	0.33	2.06		0.38	0.28	0.33	0.30	1.35
DIS	1.37	0.84	1.05	0.51	1.63	1.28	0.81	1.00	0.47	1.58	1.30	0.80	0.99	0.50	1.61	0.38	0.31	0.34	0.22	1.24
Average of (+)	1.21	0.68	0.94	0.78	1.96	1.23	0.77	0.99	0.57	1.85	1.11	0.59	0.85	16.53	2.35	0.39	0.31	0.35	0.23	1.26

Appendix 3. Statistics of the indexes. Historical volatilities, betas with respect to other indexes and correlations of the S&P 500, the Dow Jones Industrial Average Index and the Wilshire 5000 Index. Each day volatilities are calculated using 5 years of monthly data, i.e. on December 18, 2001, the volatility is the annualized standard deviation of the 60 monthly returns of the company. The returns of each month are calculated on the 18th of each month. For the periods 31/12/1996-31/5/2002, 1/1/2001-31/5/2002, and 1/12/01-31/1/02, the table contains the maximum, the minimum, the average, maximum minus minimum divided by the average, and maximum divided by the minimum.

	31/12/1996-31/5/2002						1/1/2001-31/5/2002						1/12/01-31/1/02				
	Max	min	average	(max-min)/ average	max/min		Max	min	average	(max-min)/ average	max/min		Max	min	average	(max-min)/ average	max/min
Volatility																	
S&P 500	0.20	0.07	0.14	0.89	2.62		0.20	0.14	0.17	0.30	1.37		0.19	0.17	0.18	0.14	1.15
DOW JONES IND.	0.20	0.09	0.14	0.78	2.30		0.20	0.14	0.18	0.32	1.41		0.19	0.17	0.18	0.13	1.14
WILSHIRE 5000	0.20	0.08	0.14	0.88	2.58		0.20	0.15	0.18	0.28	1.34		0.20	0.17	0.19	0.15	1.16
Beta S&P 500																	
DOW JONES IND.	1.08	0.79	0.96	0.30	1.37		0.98	0.79	0.91	0.21	1.24		0.97	0.85	0.92	0.13	1.14
WILSHIRE 5000	1.06	0.92	1.00	0.14	1.15		1.06	0.95	1.01	0.11	1.12		1.05	0.98	1.02	0.07	1.07
Beta DJ IND																	
S&P 500	0.99	0.72	0.86	0.31	1.37		0.93	0.80	0.88	0.15	1.17		0.92	0.82	0.88	0.11	1.12
WILSHIRE 5000	1.01	0.71	0.84	0.36	1.42		0.95	0.74	0.86	0.25	1.29		0.94	0.78	0.87	0.19	1.21
Beta W 5000																	
S&P 500	1.04	0.89	0.95	0.16	1.17		0.98	0.89	0.94	0.10	1.10		0.97	0.91	0.94	0.06	1.07
DOW JONES IND.	1.08	0.70	0.90	0.43	1.55		0.90	0.70	0.82	0.24	1.28		0.89	0.78	0.84	0.13	1.14
correlation																	
S&P 500 -DJ IND	0.96	0.82	0.91	0.15	1.17		0.93	0.82	0.89	0.12	1.13		0.93	0.88	0.90	0.06	1.06
S&P 500 -W 5000	0.99	0.95	0.98	0.04	1.04		0.99	0.96	0.97	0.03	1.03		0.99	0.96	0.98	0.02	1.02
DJ IND -W 5000	0.93	0.76	0.87	0.20	1.23		0.89	0.76	0.84	0.15	1.16		0.89	0.81	0.85	0.09	1.09

References

Avramov, D. (2002), "Stock Return Predictability and Model Uncertainty," Journal of Financial Economics, Vol. 64, no. 3.

Avramov, D. and T. Chordia (2001), "Characteristic Scaled Betas," SSRN Working Paper no. 276654.

Banz, R.W. (1981), "The Relationship Between Return and market Value of Common Stocks," *Journal of Financial Economics* 9, pp. 3-18.
 Basu, S. (1977), "Investment Performance of Common Stocks in Relation to their Price-Earnings Ratios: A Test of Efficient Market Hypothesis," *Journal of Finance* 32, pp. 663-682.

Bruner, R. F., K. Eades, R. Harris and R. Higgins (1998), "Best Practices in Estimating the Cost of Capital: Survey and Synthesis," *Financial Practice and Education*, Vol. 8(1), pp. 13-28.

Carelli, J.P, P. Fernandez, I. Fernandez Acin and A. Ortiz (2014), "Which is the right "Market Beta"?", http://ssrn.com/abstract=2509849

Constantinides, G. (1982), "Intertemporal Asset Pricing with Heterogeneous Consumers and Without Demand Aggregation," *Journal of Business*, vol. 55, pp. 253-267.

Copeland, T. E., T. Koller and J. Murrin (2000), *Valuation: Measuring and Managing the Value of Companies.* 3rd Edition. New York: Wiley. Damodaran, Aswath (1994), *Damodaran on Valuation*. New York: John Wiley and Sons.

Damodaran, Aswath (2001), The Dark Side of Valuation. New York: Prentice-Hall.

Elsas, R., M. El-Shaer and E. Theissen (2000), "Beta and Returns Revisited: Evidence from the German Stock Market," SSRN Working Paper no. 199428

Fama, E. F. and K. R. French (1992), "The Cross-Section of Expected Stock Returns," Journal of Finance 47, pp. 427-466.

Fernandez, P. (2015a), "CAPM: An Absurd Model", <u>http://ssrn.com/abstract=2505597</u>

Fernandez, P. (2015b), "CAPM: The Model and 305 Comments About It", <u>http://ssrn.com/abstract=2523870</u>

Griffin, J. M. (2002), "Are the Fama and French Factors Global or Country-Specific?" Review of Financial Studies, forthcoming.

- Keim, D. B. (1983), "Size Related Anomalies and Stock Return Seasonality: Further Empirical Evidence," *Journal of Financial Economics* 12, pp. 13-32.
- Keim, D. B. (1985), "Dividend Yields and Stock Returns: Implications of Abnormal January Returns," *Journal of Financial Economics* 14, pp. 473-489.
- Kothary, S. P., Jay Shanken and Richard G. Sloan (1995), "Another look at the Cross-section of Expected Stock Returns," *Journal of Finance* 50, pp. 185-224.
- Lakonishok, J. and A. Shapiro (1984), "Stock Returns, Beta, Variance and Size: An Empirical Analysis," *Financial Analysts Journal* 40, pp. 36-41.

Lakonishok, J., A. Shleifer and R. W. Vishny (1994), "Contrarian investment, Extrapolation, and Risk" *Journal of Finance* 49, pp. 1541-1578.

Lakonishok, J. and A. Shapiro (1986), "Systematic Risk, Total Risk and Size as Determinants of Stock Market Returns," *Journal of Banking and Finance* 10, pp. 115-132.

Lintner, J. (1965), "The Valuation of Risk Assets and the Selection of Risky Investments in Stock Portfolios and Capital Budgets," *Review of Economics and Statistics*, vol. 47, pp. 13-37.

Litzenberger, R. H and K. Ramaswamy (1979), "The Effects of Personal Taxes and Dividends on Capital Asset Prices: Theory and Empirical Evidence," *Journal of Financial Economics* 7, pp. 163-195.

Mas-Colell, A, M. D. Whinston and J. R. Green (1995), *Microeconomic Theory*, Oxford University Press.

Miller, M. H. (2000), "The History of Finance: An Eyewitness Account," Journal of Applied Corporate Finance, Vol. 13 no. 2, pp. 8-14.

Mossin, J. (1966), "Equilibrium in a Capital Asset Market," Econometrica, vol. 34, pp. 768-783.

Pastor, L. and R. F. Stambaugh (2001), "The Equity Premium and Structural Breaks," Journal of Finance 56, pp. 1207-1239.

Pettengill, G. N., Sundaram, S. and I. Mathur (1995), "The Conditional Relation Between Beta and Returns," *Journal of Financial and Quantitative Analysis*, 30, no. 1, pp. 101-116.

Porter, Michael (1992), "Capital Disadvantage: America's Failing Capital Investment System," Harvard Business Review, September-October.

Reilly, F. K., D. J. Wright, and K.C. Chan (2000), "Bond Market Volatility Compared to Stock Market Volatility," *Journal of Portfolio Management* (Fall), pp. 82-92.

Reinganum, M. R (1981), "Misspecification of Capital Asset Pricing: Empirical Anomalies Based on Earnings Yields and Market Values," Journal of Financial Economics, March, pp. 19-46.

Roll, R. (1977), "A Critique of the Asset Pricing Theory's Tests," Journal of Financial Economics, March, pp. 129-176.

Roll, R. (1981), "A Possible Explanation of the Small Firm Effect," Journal of Finance 36, pp. 879-888.

Roll, R. and S. A. Ross (1994), "On the Cross-Sectional Relation between Expected Returns and Betas," *Journal of Finance* 49, pp. 101-121.

Shalit, Haim and Sholomo Yitzhaki (2002), "Estimating Beta," Review of Quantitative Finance and Accounting, Vol. 18, no. 2.

Scholes, M. and J. Williams (1977), "Estimating Betas from Nonsynchronous Data," Journal of Financial Economics 5, pp. 309-327.

Shanken , Jay (1992), "On the Estimation of Beta-Pricing Models," The Review of Financial Studies 5 no. 1, pp. 1-33.

Sharpe, William (1964), "Capital Asset Prices: A Theory of Capital Market Equilibrium under Conditions of Risk," *Journal of Finance*, vol. 19, pp. 425-442.

Tinic, S. and R. West (1984), "Risk and Return: January vs. the Rest of the Year," Journal of Financial Economics 13, pp. 561-574.

Van Horne, J. C. (1992), Financial Management and Policy. Ninth edition. Englewood Cliffs, NJ: Prentice-Hall.

Welch, Ivo (2000), "Views of Financial Economists on the Equity Premium and on Professional Controversies," *Journal of Business*, Vol. 73, no. 4, pp. 501-537.

Zhang, L., Kogan, L. and J. F. Gomes (2001), "Equilibrium Cross-Section of Returns," SSRN Working Paper no. 264883.

Comments

The finding that individual beta estimates are very noisy is not new. For example, Miller and Scholes (1972) report that the sample average of the standard error of the beta estimates of all NYSE firms is around 0.32, as compared to the average estimated beta coefficient of 1.00. Thus, a random draw from this distribution of betas is going to produce any number between 0.36 and 1.64 ninety-five percent of the time. It is this imprecision in individual beta estimates (or the better known "errors in variables" problem) that motivated portfolio formation techniques of Black, Jensen, and Scholes (1972) and Fama and MacBeth (1973).

The main contribution of this paper is the way betas are calculated. Rather than estimating betas using 60 months of return data calculated using end of month prices and repeating the experiment every month, we calculate monthly returns using different days of the month over the two month period. The idea here is that the beta estimates should not show much variation as it is not likely that the true beta of these firms will change overnight. Although, intuitive at first, one should be very careful in interpreting the results documented using this approach. Moving from one day to another not only results in replacing the oldest in-sample data point with the newest observation, but also changes all the 60 data points as the returns are now calculated using different days of the month. This is, to an extent, like resampling a beta estimate from a distribution with mean one and standard deviation of 0.32 every different day in consideration (using the Miller and Scholes estimates). Thus, given the high level of standard errors, it is not surprising to observe a large amounts of day-to-day variation in individual beta estimates.

To put this into perspective, I implemented the same methodology in the paper over the same data period (December 2001-January 2002) to estimate 62 daily beta estimates for Merck. This is one of the firms the authors use in Figure 2. I find that the time series average of the 62 beta estimates for Merck is 0.607 and the time series average of the standard errors of these beta estimates is 0.204. Thus, given the level of standard error of the beta estimate, it is not surprising to obtain a high level of day-to-day variation in beta estimates for this specific firm. I calculate the standard deviation of the 62 time series beta estimates to be around 0.22 (more on this later).

Questions

What do you think about the CAPM? The CAPM tries to find an expected rate of return or a required rate of return? Which activities and professionals have benefit most of the usage of the CAPM? What is CAPM useful for? Do calculated betas change significantly from one day to the next? Which period of past returns should be used to calculate the beta: 1year... 5 years... 20 years? Which returns should be used to calculate the beta: daily, weekly, monthly...? Which market index should be used to calculate the beta: Industry Index, National Index, World Index...? Are calculated industry betas very stable? What do you conclude from table 1? Is there some "strange result" in table 2? Looking at figure 1, which company (AT&T, Boeing or Coca-Cola) was riskier? Looking at figure 3, what was a reasonale beta for AT&T? What do you conclude from table 9? What do you conclude from figure 9? What do you conclude from table A.1 of Fama and French (1992)? **Please define:** Industry beta Negative beta Dispersion of the calculated betas

Please define and differentiate:

Expected Equity Premium (EEP). Required Equity Premium (REP) Calculated beta. Risk of the shares of a company